
Downloaded from: http://ray.yorksj.ac.uk/id/eprint/4040/

The version presented here may differ from the published version or version of record. If you intend to cite from the work you are advised to consult the publisher's version: https://onlinelibrary.wiley.com/doi/abs/10.1111/aen.12425

Research at York St John (RaY) is an institutional repository. It supports the principles of open access by making the research outputs of the University available in digital form. Copyright of the items stored in RaY reside with the authors and/or other copyright owners. Users may access full text items free of charge, and may download a copy for private study or non-commercial research. For further reuse terms, see licence terms governing individual outputs. Institutional Repository Policy Statement

RaY
Research at the University of York St John
For more information please contact RaY at ray@yorksj.ac.uk
GHOST-BUSTING: Patch occupancy and habitat preferences of *Ocyale ghost* Jocqué & Jocqué 2017 (Araneae: Lycosidae), a single site endemic in north-western Madagascar

Brogan L. Pett 1,3  Joseph J. Bailey 2, 3. 
1 Biodiversity Inventory for Conservation, Brussels, Belgium; 2 Geography Department, York St. John University, York, United Kingdom; 3 Operation Wallacea, Lincolnshire, United Kingdom. 
Correspondence: Brogan L. Pett, ORCID: 0000-0002-0461-3715. Email: brogan.pett@outlook.com brogan@paralatierra.org

Abstract

Microendemic species are of great conservation concern due to their vulnerability to stochastic events and environmental change. Therefore, understanding the basic ecology of such organisms is integral to future efforts to conserve them. Here, we present the first ecological study of the only known population of *Ocyale ghost* (Lycosidae) – the single species within this genus in Madagascar – located at a small lake shore in the north-west of the island. We aimed to reveal spatial patterns of microhabitat selection and patch occupancy in relation to environmental covariates, including distance to lake edge, topography, and sand coverage. We found microhabitat selection to be strong in our occupancy models, with plots closer to the lake edge and with a higher proportion of sand showing the highest estimated occupancy. We suggest decreases in prey availability and changes in the physical characteristics of the substrate in relation to the distance from lake shore and sand availability (important for burrowing and temperature regulation), respectively, as potential factors behind occupancy. We discuss our findings in the context of the proportional underrepresentation of invertebrates in conservation research, which threatens the persistence of endemic and microendemic invertebrates in the face of significant global and local environmental change.

Key words: Burrowing, Conservation, Microendemic, Lycosidae, Madagascar, Microendemic, Occupancy modelling, Patch occupancy

Introduction

There is a profound lack of ecological knowledge on invertebrates, hampering conservation efforts (Cardoso *et al.* 2011; Gerlach, 2008). Many are thought to go extinct before being studied (Platnick and Raven 2013), let alone being assessed against IUCN Red List criteria (Regnier *et al.* 2015). Indeed,
extinction risk has been assessed for 66.7% of 69,788 described mammal species, yet only 1.9% of 1,159,365 described invertebrate species (IUCN 2018). Invertebrates fulfil crucial functions within ecosystems (Weisser and Siemann 2008) and they are, therefore, essential for not only the natural world, but also human livelihoods, health, and wellbeing (Beynon et al. 2015; Gallai et al. 2009). However, studies on invertebrates are proportionally lacking compared to other taxa and calls for population and range studies are widespread throughout the literature (Balmford et al. 2005; Butchart et al. 2007; Cardoso et al. 2011).

Endemism has long been used as a means of identifying areas for conservation prioritisation (e.g. Freitag and van Jaarsveld 1997) and endemic invertebrate species may have narrow range sizes and niche breadth, with a closer relationship with the environment than more widespread species (Rabitsch et al. 2016). In this sense, Madagascar’s levels of endemism are especially pronounced in a global context, its biota having long been recognised as exceptional (Mittermeier et al. 1998; Kremen et al. 2008; Vences et al. 2009) alongside 24 other global biodiversity hotspots (Myers et al. 2000). Endemism, proportionate to land area, is unparalleled (Mittermeier et al. 2011): 100% of native amphibians, 86% of macro-invertebrates, and over 90% of vascular plants are endemic (Wilme et al. 2012). Historical biodiversity estimates for Malagasy fauna are also proving to be underestimates (Vietes et al. 2009).

Microendemism is common in the country, with species often restricted to utilising habitat at fine spatial scales, especially in forests (Goodman and Benstead 2005) and areas surrounding low elevation river catchments (Wilme, Goodman and Ganzhorn 2006). Many invertebrates in Madagascar exhibit small populations and restricted ranges, with isolated habitats potentially harbouring microendemic species (Irwin et al. 2010).

Distribution patterns of specialist invertebrates have been shown to be particularly influenced by habitat quality (Poniatowski et al. 2018). Indeed, patch size and connectivity significantly increased patch occupancy in five threatened invertebrate species in Europe (Maes and Bonte, 2006). Meanwhile, the distribution of short-range endemic grasshoppers in South Africa was most affected by vegetation type (Adu-Acheampong et al. 2017). The role of fine-scale habitat characteristics is therefore emerging as an important general factor for invertebrate distributions.

Spiders have been shown to select living sites based on increased foraging efficiency; i.e. web building (Ruch et al. 2011), reduced predation (Rypstra et al. 2007), and the presence of beneficial species in the microhabitat (Sinclair et al. 2001). A specialist is more likely to survive and thrive in its optimal habitat because of a competitive advantage, however dispersal may be restricted (Bonte et al. 2006). For example, specialist Carabus species show reduced dispersal compared to sympatric generalists (Brouat et al. 2003) whilst increased habitat specificity is negatively associated with propensity for aerial dispersal in spiders (Bonte et al. 2003) due to the increased chances of landing in an unfavourable
habitat. This trade-off is significant when considering psammophilous Lycosidae because their range is likely to be limited by specific environmental conditions that define its niche, and it is therefore difficult to disperse across unsuitable areas that present environmental barriers (Bonte et al. 2006).

*Ocyale ghost* Jocque & Jocqué, 2017 (Lycosidae) is found on a very small habitat island, a sandy patch of less than 2ha, adjacent to a lake in Madagascar’s north-western dry deciduous forest. The genus contains large and striking species, yet a dearth of ecological and biological knowledge surrounding *Ocyale* exists (Jocque et al. 2017). *Ocyale ghost* is currently undergoing formal Red List assessment (BLP, unpubl. data). At present, only this single population is known to exist.

The use of occupancy models, developed by MacKenzie et al. (2006), has been shown to enable efficient monitoring of wildlife populations. Multiple taxa have been studied using these models: e.g. birds (Martin et al. 2009; Nichols et al. 2007), mammals (Hines et al. 2009; MacKenzie et al. 2005), and amphibians (MacKenzie et al. 2005; Kroll et al. 2008). Occupancy models that have been fit to invertebrate datasets are rarer. MacKenzie et al. (2005) provide one such example, demonstrating that occupancy models can be used over short time frames to study habitat characteristics for the giant weta *Deinacrida mahoenui* Gibbs 1999 (Anostostomatidae), a nationally threatened insect in New Zealand.

In this study, we aimed to: (1) assess the habitat characteristics associated with the distribution of *O. ghost*; (2) estimate the area occupied by *O. ghost*; (3) estimate the probability of detecting *O. ghost* in a given survey plot. Our null hypothesis was that *O. ghost*’s distribution and occupancy would be random across the site. We hypothesised that the relative area of sand across a plot would be a significant factor explaining the presence of *O. ghost*.

### 21 Materials and Methods

#### Study region

Our study centred around one of three sacred lakes at Matsedroy, (15°.49’S, 46°.65’E), in Boeny region, north-western Madagascar. The study site is within the Matsedroy forest fragment (15°29’23”S, 46°38’25”E), a community managed area without official protected status (Figure 1). The area is subject to illegal burning and logging for a variety of purposes (Evans et al. 2013). Our study area is within a previously defined zone of microendemism (Wilme et al. 2006). Temperatures remain relatively constant throughout the year (27.3°C) and mean monthly rainfall is highly variable (1-360mm) owing to the pronounced dry season from June to August and the height of the wet season from December to February (Rasamison et al. 2018). Our study site peaks at around 20m elevation, with the wider landscape of Mariarano reaching up to approximately 120m (calculated in GIS).
**Data collection**

Surveys assessing the presence or absence of *O. ghost* were carried out in plots of 9m$^2$ (3x3m). In total, 122 plots were sampled across the sandy site, separated by a minimum of 5m to ensure each quadrat could be counted as an independent sample. We chose to survey at a fine scale (9m$^2$) in order to increase our spatial data resolution on habitat changes across the small site (<2ha). Lines of plots were set in a systemic array, from points on the sandy patch, following a northerly direction as much as logistically possible, to allow an observable gradient of habitat variables leading away from the lake. Lines ranged from three to eleven quadrats in length depending on logistical considerations.

We conducted two surveys across each of the 122 plots in August 2018 (total number of surveys = 244), between 19:30 and 01:00, at least two hours after the sun had set. Plots were set up on the morning three days before surveys. This allowed sufficient time for the area to return to normal conditions after the initial disturbance caused by the surveyors. The lunar cycle was approximately one week from a full moon and there was no precipitation during the study period. Each plot was surveyed across two consecutive nights to minimise the chance of burrow relocation and reduce long-term disturbance at the location of this species’ only known population worldwide. Surveying each plot twice followed an optimum design set out by MacKenzie *et al.* (2005) based on the prior estimated detection and occupancy probabilities being above 0.6.

Each plot-based survey was carried out by a rotation of observer treatments, including a principal investigator, three University of Mahajanga MSc researchers, and a local guide. Survey rotation followed a simple pattern whereby when a plot was surveyed by a single observer, the following survey was carried out jointly by two other observers. This eliminated any bias in detection probability arising from experience or effort hours. Quadrats were considered complete as soon as an adult or sub-adult individual *O. ghost* was found. We defined the minimum size for a sub-adult as 40mm legspan from the tip of leg I to the tip of leg IV, thus explicitly excluding spiderlings from our assessment. We used marking cages of 40mm diameter with a plunger mechanism of equal size to gently press the individual into a flat splayed position, which made it quickly apparent whether or not the individual met our threshold for inclusion or was deemed a spiderling. Individuals were then gently lowered onto bisecting rulers for final confirmation by three other observers. We did not survey for spiderlings in our study as they are often excluded from optimum habitat by larger individuals (Aisenberg *et al.* 2011). Each survey at a given plot took approximately five minutes. Surveys were initially attempted from the border of the quadrat, looking for the eyeshine of an individual within the sample area. If eyeshine was detected from an individual clearly of sufficient size, the quadrat was deemed sampled. If an individual was not clearly an adult (i.e. unable to visually confirm it met our size threshold), it was captured, and size class assessed. If no individuals meeting the threshold were detected, methodical straight-line transects were
walked across the plot, looking for burrows (for closer inspection) or individuals at poor angles for
tapetum reflection (the main reason for initial detection failure).

Environmental variables were taken for covariate analyses from each plot. The temperature was
recorded at the beginning of the survey using a Benetech GM550 infrared thermometer. After all
quadrats were gridded out using a 20m measuring tape, a substrate assessment was carried out at each
one. Each of four observers stood at the corners of the quadrant and independently estimated the
proportion of area that was covered by: (1) bare ground (sand); (2) leaf litter (common around habitat
periphery (forest)); (3) grasses (see Appendix 1); and (4) low lying shrub. Estimates were averaged
between the four values obtained from each observer, for each variable. Prior to analysis, all plot
substrate proportions were arcsine transformed.

Topographic metrics (elevation, aspect, and slope) were derived from the Japan Aerospace Exploration
Agency (©JAXA; Takaku and Tadono 2017) 30m global digital surface model in ArcGIS 10.3. Aspect
was measured in degrees, converted to radians, then a cosine (for north-south) and sine (for east-west)
transformation was applied prior to analysis, to represent north and east as high values (1) and south
and west as low values (-1).

**Statistical analysis and model selection**

We used single season occupancy models described in MacKenzie *et al.*, 2006, which use binary
detection/non-detection data to assess the support for competing hypotheses about the distribution of
species. Occupancy models also incorporate the possibility of false absences in datasets by using
detection probabilities (notation ‘*p*’), using them to make robust inferences about wider occupancy
probabilities. When unaccounted for, imperfect detection can alter occupancy estimates to a large
degree (MacKenzie *et al.* 2006).

One of the key considerations in single-season occupancy models is meeting the assumption of
population closure during the study period. We believe the closure assumption can be reliably assumed
because: (1) our surveys took place on consecutive nights; and (2) our observation that disturbed
individuals often retreated back to a visible or previously invisible (to the observer) burrow. Three
sexually mature adult males did not retreat to an active burrow once released (*BLP personal
observation*) but analysing this (currently) statistically negligible transience behaviour in reproductive
state males is beyond the scope of this study. Burrow construction in sand-dwelling spiders is
energetically expensive (De Simone, Aisenberg and Peretti 2015), and thus the probability of an
individual changing burrow site over 24 hours is assumed to be very low.

Each set of observations at a plot generates a detection history for the target species. For example, in a
hypothetical study, the binary notation 101 denotes the target species was detected on the first and third
sampling occasions, but not on the second. Thus, the probability statement for this detection history can be expressed mathematically as:

\[ \text{Pr}(h_1=101) = \Psi p_1 (1-p_2)p_3 \]

Where ‘Pr(h_1)’ is the probability of observing the detection history gathered, in the above case 101. Detection is represented by the true occupancy (\(\Psi\)) multiplied by detection probability at the first (\(p_1\)). Non-detection is represented as 1-\(p_2\) (1- probability of detection at occasion two). Another example, the detection history 001 would signify a detection only on the last of three sample occasions, expressed as:

\[ \text{Pr}(h_1=001) = \Psi (1-p_1)(1-p_2)p_3 \]

We used a logit link function to estimate the magnitude of influence that ecological covariates (Table 1) had on the parameters of interest (\(\Psi\) and \(p\)). Logit link functions serve to take a linear combination of covariate values and convert them to a scale of probability (MacKenzie et al. 2006). Occupancy probability can be modelled with a logit link function as:

\[ \text{logit} \Psi_i = \delta_0 + \delta_1 x_1 + \delta_2 x_2 + \ldots + \delta_n x_n \]

Where \(\delta_0\) is the intercept value and \(\delta_1\) is the effect size of covariate \(x_1\).

Model selection was made using Akaike’s Information Criterion (AIC) (Akaike 1973), which can be expressed mathematically as:

\[ \text{AIC} = -2\ln[L(\Theta|x)] + 2 \delta \]

Where \(\delta\) is the number of parameters estimated in the model. The penalisation of assessing an increased number of parameters encourages parsimony (i.e. the simplest model with the highest explanatory power). AIC is a robust method of statistical inference, used widely in ecological studies. Aho et al. (2014) found that 84% of ecology studies between 1993-2013 that used multi-model inferences evaluated using AIC as opposed to other methods of evaluation. Models are evaluated by using calculated Delta AIC (\(\Delta\)) values, which can be expressed mathematically as:

\[ \text{Delta AIC}_i = \Delta_i = \text{AIC}_i - \min \text{AIC} \]

Where \(\text{AIC}_i\) is the AIC value for model \(i\), and \(\min \text{AIC}\) is the AIC value of the model with the strongest support. The highest-ranking candidate models are considered with an \(\Delta\text{AIC}\) value <2 models with \(\Delta\text{AIC} >7\) are disregarded as having virtually no support, and those in between receiving moderate support (Burnham and Anderson, 2002).

We fitted occupancy models using the ‘occu’ function in package ‘unmarked’ (Fiske and Chandler 2011) in R version 3.5.2 (R Core Team, 2013). We assessed model fit between our top ranked model
and the reference distributions by comparing the Chisq statistic, using the ‘MacKenzie and Bailey goodness of fit test’. The function ‘mb.gof.test’ in the package ‘AICcmodavg’ (Mazerolle, 2019), with 1000 simulations, showed there was no indication of a lack of fit ($P > 0.05$). MacKenzie et al. (2006) note that occupancy modelling without adequate tests of model fit risk weak inferences.

Collinearity between plot-specific variables was generally negligible (Appendix 3). In particular, our key variables in the final models were largely independent of one another and aspect and slope were only weakly correlated or had no significant correlation. For example, collinearity was low between distance to lake edge and east-west aspect ($r = 0.21; p < 0.05$) and bare ground ($r = -0.06; p > 0.05$).

All such variables were therefore analysed as separate factors. However, as expected, the proportion of short grass showed high collinearity with proportion of bare ground ($r = -0.785; p < 0.05$), which was considered when interpreting the results.

**Results**

We found that microhabitat selection was strong in *O. ghost*. The most important and significant plot-scale covariate predicting occupancy was distance to water, with plots closer to the lake’s edge exhibiting markedly higher occupancy. The proportion of bare sand in a plot was also a significant predictor, with the effect inversed; plots with little sandy substrate had significantly lower occupancy probabilities. Detection was not related to the temperature at the time of survey, however temperature variation during our survey period was relatively low.

**Probability of occupancy in relation to habitat covariates**

The most parsimonious model based on AIC weights was model “Ψ (DW+BGr+AspEW) P (.)” ($w = 0.65$, Table 2). This model indicated that occupancy of *O. ghost* was related to the plot’s distance to the water’s edge, the proportion of bare ground (sand), and the east-west aspect (increased occurrence on west-facing slopes) of the plot. Our second most highly ranked model was structured the same but for the removal of ‘AspEW’ as a covariate “Ψ (DW+BGr) P (.)” ($w = 0.19$). We included the null model “Ψ (.) P (.)” in our final table of results for comparison, receiving no statistical support ($w = <0.0001$); we can firmly reject the null hypothesis that habitat covariates within plots do not influence *O. ghost* distribution.

Using the coefficient values estimated from our top model (Table 3), we show that distance to water source (‘DW’ in model notation) has a strong negative contribution to the probability of occupancy ($Z$ value = -3.80, $P <0.001$) (Figure 2). Our hypothesis that the proportion of bare ground would have a strong positive relationship with *O. ghost* occupancy was supported by our results, ($Z$ value = 2.41, $P = 0.016$), although this effect was somewhat weaker than distance to water.
Our models were chosen to represent a diversity of plausible scenarios that could affect *O. ghost* occupancy and ascertain which covariates or covariate pairs influenced occupancy. To avoid data dredging (Burnham and Anderson, 2002), we chose eleven models *a priori*. Our final top ranked model was constructed as a 12th due to its anticipated explanatory power based on initial results. This final model’s residuals showed very low spatial autocorrelation (Moran’s I = 0.118; p > 0.05), supporting covariate selection.

*Probability of occupancy*

Traditional presence-absence methods (naïve occupancy) would estimate the survey site as having 55% of plots occupied (67/121). However, using model averaged estimates of true occupancy, by estimating the detection parameter, we estimate that 63% (SE 7%, Table 4) of plots were occupied (76/121). This suggests that nine plots were occupied by target sized individuals of *O. ghost* and not detected by observers.

*Probability of detection*

We initially wanted to evaluate any effect of temperature on the detection probability of *O. ghost*. However, when modelled, models with detection held constant \( P(.) \) were substantially more supported than any individual model where constant detection was substituted for temperature \( P(temp) \) or time variation \( P(t) \). Thus, we excluded these models from our final analysis. The model averaged detection estimate was 0.74 (SE= 0.05) (Appendix 2), validating our initial assertion during survey design that detection probability would be above 0.6 (see Methods). However, our observed detection probability was lower than predicted.

**Discussion**

We have found strong evidence that microhabitat selection is a key factor influencing the distribution of *O. ghost*, following our assessment of fine-scale habitat characteristics. Distance to the lake edge and proportion of bare ground appeared to be key factors for this microendemic spider (Aim 1). Indeed, occupancy estimates were dramatically reduced at greater distances from the water (i.e. the edge of the lake) (2). Our top-ranking model estimated occupancy at ≥0.85 for plots within 30 metres of the lake edge, dropping to 0.5 at just under 100 metres; at >150m from the lake edge, estimated occupancy reduced substantially to 0.2. Our most parsimonious model “Ψ (DW+BGr+AspEW) P(.)” accounted for 0.65 of AICc model weight, with a single variation our second ranked model accounted for 0.19 model weight (the exclusion of aspect east-west). Our top ranked model also showed very low spatial autocorrelation in the residuals, suggesting key facets of the environment were adequately accounted for. Combined, our two best models carry over 80% of AICc weight, the only variation in model structure being the removal of east-west aspect in the second favoured model. The probability of
detecting *O. ghost* in a plot was relatively high (3), but possibly reduced by the lower observed

- temperatures, which would have influenced spider activity levels.

Results were consistent across the models. Those incorporating both distance to the water’s edge and

- the proportion of bare ground within a plot had over 92% support, demonstrating a clear relationship

- between these variables and *O. ghost* occupancy. Insect biomass can be higher closer to water sources

- (and therefore at the lower elevation), which has been shown to increase resource utilisation by

- predators such as Lycosidae spiders (Gratton, Donaldson and Zanden 2008), as well as increase the

- overall abundance and biomass of spiders (Henschel, Mahsberg and Stumpf, 2001). Alternatively, the

- area surrounding the lake could support a microclimate that is more stable than those at the periphery

- of the habitat; for example, it might be less susceptible to weather extremes (Lubin, Henschel and Baker

- 2001).

The proportion of bare ground (sand) in a plot was hypothesised to increase the occupancy of *O. ghost*,

- chiefly due to the species’ psammophilous specialisation. A specialised species will outcompete a

- generalist in optimal habitat (Bonte *et al.* 2006), corroborating our observations that other Lycosidae

- species are at very low abundances on the sandy habitat (*BLP personal observation*). Sand is a critical

- substrate for *O. ghost* burrow construction, it being less energy expensive to construct burrows than in

- soil (Suter, Stratton, and Miller 2011). Burrowing Lycosidae species have also been shown to be highly

- tenacious and venture only short distances from burrows (Suter, Stratton, and Miller, 2011). However,

- the substrate itself may be influenced by the lake, which may cause moisture gradients that affect

- individual spiders’ ability to burrow and forage. Therefore, the prominence of lake distance in our

- results may be moderated by its effect on the substrate, but further work on these abiotic factors, and

- their interaction, is needed.

Our results show an association between *O. ghost* occupancy and the aspect of the slope, with a

- preference for west-facing slopes. Some Lycosidae species have been shown to prefer certain sides and

- slopes of sand dunes for their burrow site (Aisenberg *et al.* 2011). The same study also showed a clear

- difference between burrow site location and age class of Lycosidae spiders. Territorial exclusion and

- intraguild predation have been demonstrated in burrowing Lycosidae species previously (Moya- Lараño

- *et al.* 2002), and it is a likely cause of age class separation. Sex is known to be a key factor for some

- wolf spider activity patterns (Framenau, 2005; Cera and Spuńgis, 2011). Whilst this was beyond the

- scope of our present study, it warrants further investigation, given that burrowing Lycosidae may show

- unusual patterns in relation to sex, such as reversed sex roles (Aisenberg and Costa, 2008).

The preference for west-facing slopes may simply be because this is where most of the suitable habitat

- was in our study area, it may relate to insolation, or be due to the interaction of these factors. West-

- facing slopes are likely to have higher insolation than east-facing slopes, potentially increasing the

- average and maximum daily temperature of burrows, depending on local topography. Predatory
arthropods living in sand are exposed to higher climate extremes, with hot days and substantially cooler nights (Bonte and Maelfait 2001) and these fine-scale temperatures are likely to be affected by east-west aspect. Maintaining a relatively constant burrow temperature may reduce extreme metabolic fluxes, increasing individual survival. Longer burrows offer a buffer to external heat fluctuations in Lycosidae (Aisenberg et al. 2011), but we did not measure burrow temperature, and thus our inferences about the role of aspect is speculative.

Burrow site selection in a species of Eresidae spider in Namibia was suggested as having been influenced by conspecific cuing (Birkhofer et al. 2012), potentially buffering the population against extreme climatic events. Meanwhile, patch quality was a significant predictor of occurrence in another Lycosidae species (Bonte et al. 2003). A species of Pardosa (Lycosidae) was shown to select substrate with less variable temperatures in winter (Kraus and Morse 2005) and a web-building Oecobiidae spider was shown to select strongly for substrate during web construction (Voss et al. 2007). Despite the dearth of ecological studies on habitat selection in spiders, studies that deal with the topic clearly report consistent results.

We are confident in our data from the few sampling occasions. Utilising such ‘snap- shot’ data has been effectively demonstrated elsewhere, following optimum design set out by MacKenzie et al. 2006, e.g.: toads (Bradford et al. 2003), hemipterans (Biedermann, 2004); and orthopterans (MacKenzie et al. 2005). The fact that we were in the field for a single season does pose its own limitations. Temporally, conducting a similar study during the wet season would elucidate whether or not these patterns of spatial distribution are consistent. Curtis and Bloch (2014) showed that microhabitat selection shifted seasonally in an Amblypygid, another terrestrial predatory arthropod. Thus, the question of O. ghost’s wet season habits emerges. The observed night-time temperatures during the study period were relatively low (average 19.8°C) and having an equal sampling period during the same season with greater temperature variation could prove insightful. However, due to the low likelihood of regular burrow relocation (see Methods), we believe the temperature had minimal influence on the observed distribution of O. ghost. Temperature could influence the activity patterns and hence detection probabilities, though, as demonstrated in other ectothermic invertebrates (Høye and Forchhammer 2008), including spiders (Pruitt, Demes and Dittrich-Reed 2011).

Goodman and Benstead (2005) showed that of 5,808 species of non-marine invertebrate in Madagascar, 4,976 (86%) of species are endemic. They also classified the spider fauna as ‘relatively well known’ at 459 species (pp. 74). However, in the years since, numerous publications have described dozens of new species at a time (e.g. Álvarez-Padilla et al. 2012; Wood 2008; Wood and Scharff 2018). Some research even includes currently undescribed species (e.g.: Henrard and Jocque, 2016; Bauer et al. 2018). Such is the complexity of Madagascar’s spider fauna.
Madagascar’s biodiversity is exceptional, largely as a result of the endemism and microendemism across multiple taxa (e.g. Wilme et al. 2006). *Neogrosphus* scorpions (Buthidae) contain only three species, yet exhibit species ranging from the extreme subarid conditions of the south-west, to the subhumid evergreen forests of the north-east (Lourance et al. 2015). Isolated populations of *N. griveaudi* (Vachon 1969) exist more than 300km apart, a likely prequel to allopatry. Species with low dispersal capabilities (e.g. scorpions) in a country characterised by processes capable of fragmenting habitat areas into small isolated islands (see Lourenco et al. 2015 for a detailed treatment) are a chief component of the high species richness of Madagascar (Mittermeier et al. 2011). Microendemism is not limited to arachnids, however (e.g. *Brookesia* chameleons (Townsend et al. 2009); Microhylid frogs (Scherz et al. 2019)).

Only 238 of 47,904 described spider species have been assessed against IUCN Red List criteria (IUCN 2018; World Spider Catalog 2018), with the description of new species showing no signs of slowing (Platnick and Raven 2013). A growing body of work demonstrates the paucity of invertebrate conservation research compared to other organisms (Clark 2002), including the megadiverse order of Araneae. Indeed, when species richness is corrected for, arthropods are heavily selected against as research study organisms (Rosenthal et al. 2017). Cardoso et al. (2011), note that basic science and funding is lacking for invertebrates, a key factor impeding invertebrate conservation. Consequently, a lack of knowledge is one of the key factors threatening invertebrates, their associated ecosystem functions, and species that depend upon them.

Habitat destruction, fragmentation, and degradation are key drivers of species endangerment in Madagascar, and indeed globally (Harper et al. 2007). Many of the spider species assessed as endangered or critically endangered cite these factors as a cause; e.g. *Anapistula ataecina* (Cardoso 2010), *Hogna ingens* (Cardoso 2014), *Mariblemma pandani* (Gerlach 2014). In the case of *Ocyale ghost*, its sandy habitat is relatively safe from the more common forms of habitat loss because the area is a designated sacred lake by the Sakalava people living in neighbouring villages. However, the stability of sand dynamics (the stability of the habitat) could be affected by deforestation. For example, deforestation in the surrounding area decreases soil quality (Islam and Weil 2000), which reduces soil stability, and could allow an increase in sand dynamics. The area is also susceptible to increasing disturbance as the human population in the surrounding villages increases, with livestock utilising the nearby lake daily in this dry landscape (*pers. obs.*). Increased cattle grazing decreases occupancy of sand dwelling invertebrates (Bonte and Maes 2008). Although, cattle grazing reduces vegetation cover significantly, which has been shown to alter the spider community by way of suppressing guilds and species that utilise vegetation (Warui 2004). Therefore, in a relatively unique biological case, the overall net impact of cattle grazing and deforestation on *O. ghost* is unknown. Unfortunately, however, due to the very small habitat area, stochastic vulnerability to extinction is likely to be high for this species.
Compounding these localised threats, climatic change in the area is likely to create drier, warmer conditions, affecting water availability and ground temperatures (Hannah et al. 2008). These factors will undoubtedly increase the frequency of droughts and the severity of fires, which could increase mortality of individual spiders. Insect species have been shown to shift upwards in elevation to combat the increased temperature at lower elevations (Wilson et al. 2007). However, this is unlikely in our study species due to its low elevation lake shore habitat and the unsuitable surrounding habitat matrix. *O. ghost* also favoured plots closer to the water source: as the climate dries, the water level of the lake and surrounding water sources is likely to decrease, further augmenting the loss of suitable habitat (if such sandy plots persist at all).

Madagascar is changing dramatically, with land use and climate change predicted to significantly alter habitats for many taxa, as well as overall biodiversity and ecosystem function (Thuiller et al. 2006; Schatz et al. 2008). Narrow-range endemics and habitat specialists are placed at a substantial evolutionary disadvantage because of their lower dispersal capabilities and inability to move to suitable habitat (assuming it exists, and it is accessible to the species). As a result of this, endemic and threatened species are, inherently, at the greatest risk of extinction and of greater conservation concern and public interest (Williams et al. 2000). It has also been demonstrated elsewhere that, nationally, areas containing rare and endemic species may highlight areas of higher conservation priority (e.g. Bonn et al. 2002). Borders of distinct habitat types also create various degrees of edge habitat, increasing activity of some invertebrates (De Smedt et al. 2019) and potentially allowing some species to exist in the ‘transition’ zone (Brown and Hutchings 1997).

**Conclusion**

Our single-season occupancy models provided strong evidence that microhabitat selection is strong in *O. ghost*. Our null model was rejected and key habitat covariates were elucidated. The species was not detected at any of fourteen other potentially suitable sites in the region (BLP, unpubl. data). Thus, assessing the ability and usage of aeronautic ballooning dispersal, or dispersal of any kind, in *O. ghost* should be central to future study of the species. Without such a mechanism, it is likely this is indeed the only population of *O. ghost* and, as such, we work towards listing the species in one of the IUCN ‘endangered’ categories. This species is the only known representative of its genus in Madagascar and may represent a distinct evolutionary lineage. Building up a knowledge base on the fine-scale ecology of likely threatened species is vital towards future efforts to conserve them. Conservation of this charismatic species would be inexpensive and have benefits for other species; it is therefore crucial to the integrity of conservation efforts in this unique part of the world.
1 Acknowledgements
This work was supported by the Mohamed bin-Zayed Species Conservation Fund (project number 180518182) and the Percy Sladen Memorial Fund. Our thanks go to the Ministry of Forests and Environment for providing permits (№ 14C/18/MEEF/GFF/DGF/DSAP/SCB.Re) and permissions to work in the region. We also thank Operation Wallacea and Development and Biodiversity Conservation Action for Madagascar (DBCAM) for logistical support and resources around the fieldwork. The people of Mariarano kindly offered support and help with lodgings and staffing. BLP thanks University of Mahajanga MSc students: Paul Bienvenue Rabemananjara, Fifaliana Andriamiravo, Romario Nasondrotra, for assistance in the field. BLP also thanks Dr. Merlijn Jocqué, who offered instrumental initial advice. JJB thanks York St John University for their continued support.

11 References


De Smedt, P., Baeten, L., Proesmans, W., Van de Poel, S., Van Keer, J., and Giffard, B. et al. (2019). Strength of forest edge effects on litter-dwelling macro-arthropods across Europe is influenced by forest age and edge properties. Diversity and Distributions. doi: 10.1111/ddi.12909


Tables

Table 1. Predictors (and abbreviations) used in the modelling.

<table>
<thead>
<tr>
<th>Predictors used in logit link (abbreviation)</th>
<th>Description</th>
</tr>
</thead>
<tbody>
<tr>
<td>Bare ground (BGr)</td>
<td>Bare sand substrate without stones, leaf litter or other objects.</td>
</tr>
<tr>
<td>Distance to water (DW)</td>
<td>Distance between centroid of plot and the nearest edge of the adjacent lake.</td>
</tr>
<tr>
<td>Short grass (ShGr)</td>
<td>Below knee-height grass that grows in concentrated patches at site (see Appendix 1).</td>
</tr>
<tr>
<td>Leaf litter (LeLit)</td>
<td>Leaf litter often at the periphery of the site, blown from nearby forest. Also included patches of broken branches that created an inconsistent substrate area.</td>
</tr>
<tr>
<td>Low shrub (L.Shr)</td>
<td>Areas of dry shrubby vegetation (not grasses) that generally appear as sporadic and sparse.</td>
</tr>
<tr>
<td>Elevation (Elev)</td>
<td>Metres above sea level.</td>
</tr>
<tr>
<td>Slope (Slp)</td>
<td>Topographic slope in degrees. Calculated in ArcGIS.</td>
</tr>
<tr>
<td>Aspect EW (AspEW)</td>
<td>The degree to which the slope faces east or west. Calculated in ArcGIS.</td>
</tr>
<tr>
<td>Aspect NS (AspNS)</td>
<td>The degree to which the slope faces north or south. Calculated in ArcGIS.</td>
</tr>
</tbody>
</table>

Table 2. Model comparisons to identify ecological patch covariates that influence ghost spider occupancy using single season survey design.

<table>
<thead>
<tr>
<th>Model</th>
<th>AIC</th>
<th>ΔAIC</th>
<th>AIC weight</th>
<th>Cumulative Weight</th>
<th>K*</th>
</tr>
</thead>
<tbody>
<tr>
<td>$\Psi (DW+BGr+AspEW) P (.)$</td>
<td>260.26</td>
<td>0.00</td>
<td>0.65</td>
<td>0.65</td>
<td>5</td>
</tr>
<tr>
<td>$\Psi (DW+BGr) P (.)$</td>
<td>262.65</td>
<td>2.39</td>
<td>0.19</td>
<td>0.84</td>
<td>4</td>
</tr>
<tr>
<td>$\Psi (DW+BGr+Slp) P (.)$</td>
<td>264.46</td>
<td>4.20</td>
<td>0.08</td>
<td>0.92</td>
<td>5</td>
</tr>
<tr>
<td>$\Psi (DW+L.Shr) P (.)$</td>
<td>264.85</td>
<td>4.59</td>
<td>0.07</td>
<td>0.99</td>
<td>4</td>
</tr>
<tr>
<td>$\Psi (.) P (.)$</td>
<td>298.12</td>
<td>37.85</td>
<td>&lt;0.0001</td>
<td>&lt;0.0001</td>
<td>2</td>
</tr>
</tbody>
</table>

Notes: K= number of parameters, DW= Distance to water (m), BGr= Bare ground, AspEW= Aspect East/West, Slp= Slope, L.Shr= Low shrubs, (.) denotes a parameter held constant. See methods section for treatment of covariates and statistical summaries.

Table 3. Coefficients for our top ranked model “$\Psi (DW+BGr+AspEW) P (.)$” (see table 2)

<table>
<thead>
<tr>
<th>Estimate</th>
<th>SE</th>
<th>Z-value</th>
<th>P</th>
</tr>
</thead>
</table>
Table 4 Occupancy estimates for each model from Table 1

Note: model estimates were obtained by fixing covariates at their mean values

<table>
<thead>
<tr>
<th>Model</th>
<th>( \Psi )</th>
<th>SE</th>
</tr>
</thead>
<tbody>
<tr>
<td>( \Psi (DW+BGr+AspEW) P (.) )</td>
<td>0.69</td>
<td>0.09</td>
</tr>
<tr>
<td>( \Psi (DW+BGr) P (.) )</td>
<td>0.62</td>
<td>0.07</td>
</tr>
<tr>
<td>( \Psi (DW+BGr+Slp) P (.) )</td>
<td>0.65</td>
<td>0.10</td>
</tr>
<tr>
<td>( \Psi (DW+L.Shr) P (.) )</td>
<td>0.61</td>
<td>0.06</td>
</tr>
<tr>
<td>( \Psi (.) P (.) )</td>
<td>0.59</td>
<td>0.05</td>
</tr>
<tr>
<td>Model averaged</td>
<td><strong>0.63</strong></td>
<td><strong>0.07</strong></td>
</tr>
</tbody>
</table>
Figures

Fig 1 (a) The study’s location in the north-western dry forest within Madagascar, also showing elevation (max = 2744m); (b) Our study lake within the landscape, also showing nearby lakes to the north and
north-east; (c) our study lake in isolation, showing sampled points on the sandy northern/north-eastern shore and a connected sandy area to the east, where blue indicates a presence and light green an absence. All satellite imagery is from SENTINEL (https://www.copernicus.eu/en), July 2018, and the elevation data in (a) is from WorldClim (Hijmans et al. 2005).

**Fig 2** Estimated probability of occurrence (grey shading shows upper and lower 95% confidence intervals) in relation to distance from the lake shore.